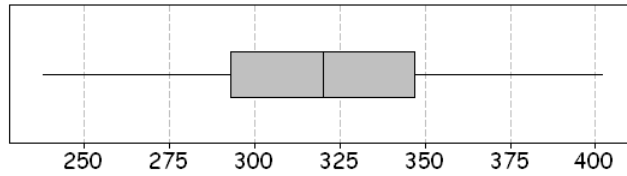


Answers to November 2008 exam

1. (1-1) D; [$0.6 \times 0.02 + 0.3 \times 0.03 + 0.1 \times 0.04 = 0.025$]
 (1-2) B; [$\frac{0.004}{0.025} = 0.16$]
 (1-3) E; [$p_{22}^{(2)} = 0.33$]
 (1-4) E; [$\text{var}(Z) = 4^2 + 3^2$]
 (1-5) C; [the Normal distribution has shorter tails than exponential; it must therefore have an increasing hazard function.]
 (1-6) A; [$(Z > 1) \cap (Z > 2) = (Z > 2)$, since $(Z > 2) \subset (Z > 1)$]
 (1-7) C; [by definition]
 (1-8) B; [$z = \frac{15.5-15}{0.204} = 2.45 \Rightarrow p = 0.014$]
 (1-9) D; [$\text{RLL} \geq -2 \Rightarrow -3 < \theta < 15$]
 (1-10) A; [$df = (3, 3, 3, 6, 15)$; $F = \frac{40}{10} = 4$, cf. $F_{3,6}$]
2. (a) i. $\Pr(F \cup G) = \Pr(F) + \Pr(G) \Rightarrow 0.8 = 0.2 + g \Rightarrow g = 0.6$;
 ii. $\Pr(F \cup G) = \Pr(F) + \Pr(G) - \Pr(F \cap G) \Rightarrow 0.8 = 0.2 + g - 0.2g \Rightarrow g = 0.75$;
 iii. probability table:
- | | | | |
|------|---------|---------|-----|
| | G | G' | |
| F | $g-0.6$ | $0.8-g$ | 0.2 |
| F' | 0.6 | 0.2 | 0.8 |
| | g | $1-g$ | 1 |
- $\Pr(F|G) = \frac{g-0.6}{g}$; and $g \leq 0.8$, so $\Pr(F|G) \leq \frac{0.2}{0.8} = 0.25$.
- (b) $X \stackrel{d}{=} \text{Bi}(n=20, p=0.1)$, so $\Pr(X \leq 2) = 0.1216 + 0.2702 + 0.2852 = 0.677$ (Tables).
- (c) i. $P = \begin{bmatrix} 0.2 & 0.8 & 0 & 0 \\ 0.1 & 0.1 & 0.8 & 0 \\ 0.1 & 0 & 0.1 & 0.8 \\ 0 & 0 & 0 & 1 \end{bmatrix}$;
- ii. five-step transition probabilities are given by P^5 ; $p_{14}^{(5)}$, the (1,4)-element of P^5 , gives the probability of going from state 1 to state 4 in five steps, i.e. the probability of completion in 5 hours.
3. (a) $E(X) = 0 \times 0.2 + 1 \times 0.6 + 2 \times 0.2 = 1$;
 $E(X^2) = 0 \times 0.2 + 1 \times 0.6 + 4 \times 0.2 = 1.4 \Rightarrow \text{var}(X) = 0.4$.
 $\therefore E(Y) = 160 \times 1 = 160$, $\text{var}(Y) = 160 \times 0.4 = 64$.
 $\Pr(Y \leq 150) \approx \Pr(Y^* < 150.5) = \Pr(Y_s^* < -1.188) = 0.118$.
- (b) $F(t) = 1 - e^{-H(t)}$, where $H(t)$ denotes the cumulative hazard function.
 $H(4) = \int_0^{0.4} 0.03\sqrt{t} dt = [0.03 \frac{4^{3/2}}{3/2}] = 0.16$;
 $\therefore \Pr(T > 4) = e^{-H(4)} = e^{-0.16} = 0.852$.
- (c) approx formula with $\psi(x) = \sqrt{x}$, so $\psi'(x) = \frac{1}{2\sqrt{x}}$, $\psi''(x) = -\frac{1}{4x\sqrt{x}}$; $\mu = \theta$, $\sigma^2 = c\theta^2$;
 $E(\sqrt{Z}) \approx \sqrt{\theta} + \frac{1}{2}(-\frac{1}{4\theta\sqrt{\theta}})c\theta^2 = \sqrt{\theta} - \frac{1}{8}c\sqrt{\theta} = \sqrt{\theta}(1 - \frac{c}{8})$;
 $\text{var}(\sqrt{Z}) \approx (\frac{1}{2\sqrt{\theta}})^2 c\theta^2 = \frac{c\theta}{4}$.
 The approximations are best when $\text{var}(Z)$ is small, i.e. if c is small.
- (d) i. $\Pr(X > 50) = \Pr(X_s > 1) = 0.159$;
 ii. $X - T \stackrel{d}{=} N(-4, 5^2)$, since $46 - 50 = -4$ and $3^2 + 4^2 = 5^2$.
 $\Pr(X > T) = \Pr(X - T > 0) = \Pr((X - T)_s > 0.8) = 0.212$.
4. (a) $\bar{x} = 27.14$, $s = 9.85$;
 (b) $\min \sim c_{0.02} = 320 - 2.0537 \times 40 = 237.9$;
 $Q1 \sim c_{0.25} = 320 - 0.6745 \times 40 = 293.0$;
 $\text{med} \sim c_{0.5} = 320$;
 $Q3 \sim c_{0.75} = 320 + 0.6745 \times 40 = 347.0$;
 $\max \sim c_{0.98} = 320 + 2.0537 \times 40 = 402.1$;

five-number summary: (238, 293, 320, 347, 402).



(c) “ x -coordinate” = $\Phi^{-1}(\frac{15}{20}) = 0.6745$; “ y -coordinate” = $x_{(15)} = 65$.
 y -intercept $\Rightarrow \hat{\mu} = 50$; slope $\Rightarrow \hat{\sigma} = 20$ ($= \frac{70-50}{1-0}$)

5. $\bar{X} \stackrel{d}{=} N(\mu, \frac{5^2}{12})$

i. 95% probability interval for \bar{X} : $40 \pm 1.96 \times \frac{5}{\sqrt{12}} = (40 \pm 2.83) = (37.17, 42.83)$.

ii. size = $\Pr(|\bar{X} - 40| > 3)$, where $\bar{X} \stackrel{d}{=} N(40, \frac{5^2}{12})$;

$$\therefore \text{size} = 2\Pr(\bar{X}_s > \frac{3\sqrt{12}}{5}) = 2\Pr(\bar{X}_s > 2.078) = 0.038.$$

iii. $p = 2\Pr(\bar{X} > 44.0)$, where $\bar{X} \stackrel{d}{=} N(40, \frac{5^2}{12})$;

$$\therefore p = 2\Pr(\bar{X}_s > \frac{4\sqrt{12}}{5}) = 2\Pr(\bar{X}_s > 2.771) = 0.006.$$

iv. power = $\Pr(|\bar{X}' - 40| > 3)$, where $\bar{X}' \stackrel{d}{=} N(45, \frac{5^2}{12})$;

$$\therefore \text{power} = 1 - \Pr(37 < \bar{X}' < 43) = 1 - \Pr(-5.543 < \bar{X}'_s < -1.386) = 0.917.$$

v. 95% confidence interval for μ : $44.0 \pm 1.96 \times \frac{5}{\sqrt{12}} = (44.0 \pm 2.83) = (41.17, 46.83)$.

vi. 95% prediction interval for X : $44.0 \pm 1.96 \times 5\sqrt{\frac{13}{12}} = (44.0 \pm 10.2) = (33.8, 54.2)$.

6. (a) $\hat{\lambda} = \bar{x} = 5.7$; $\sigma^2 = \text{var}(X) = \lambda^2$, so $\text{var}(\bar{X}) = \frac{\lambda^2}{10}$;

$$\text{se}(\hat{\lambda}) = \frac{\hat{\lambda}}{\sqrt{10}} = \frac{5.7}{\sqrt{10}} = 1.8.$$

(b) $\frac{\partial \ln L}{\partial \theta} = -20 + \frac{100}{\theta} = 0 \Rightarrow \hat{\theta} = 5$;

$$-\frac{\partial^2 \ln L}{\partial \theta^2} = \frac{100}{\theta^2} \Rightarrow \text{se} = 1/\sqrt{\frac{100}{25}} = 0.5.$$

(c) i. tables: $E(Z_{(12)}) = 1.629 \Rightarrow E(R) = 2 \times 1.629\sigma$;

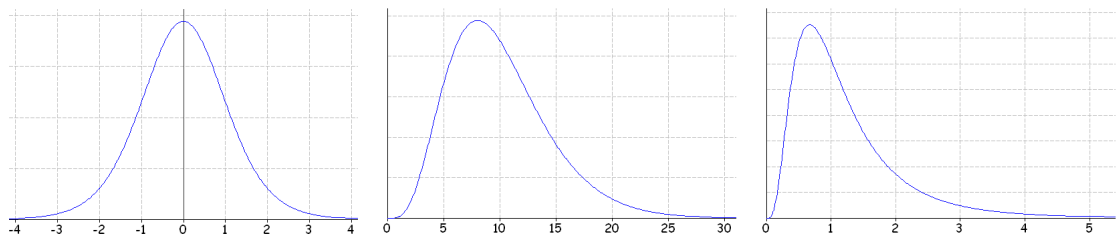
$$\hat{\sigma} = \frac{\bar{R}}{3.258} = \frac{7.32}{3.258} = 2.25.$$

ii. control limits for \bar{X} : $36.2 \pm 3 \times \frac{2.25}{\sqrt{12}} = (36.29 \pm 1.95) = (34.34, 38.24)$.

iii. tables: $E(Z_{(11)}) = 1.116 \Rightarrow E(Q) = 2 \times 1.116\sigma$; thus $c = 2.232$.

The estimator $\hat{\sigma}^* = \frac{\bar{Q}}{2.232}$ is more stable than $\hat{\sigma} = \frac{\bar{R}}{3.258}$. It is less susceptible to outliers, which is both good and bad. It may be important to pick up outliers; but it is important to have a good estimate of the underlying variation.

7. (a) t_{10} pdf χ_{10}^2 pdf $F_{10,10}$ pdf



(b) i. tables: $\Pr(0.226 < F_{8,8} < 4.433) = 0.95$,

$$95\% \text{ CI for } \frac{\sigma_1^2}{\sigma_2^2}: (0.226 \times \frac{7.41^2}{5.67^2} < \frac{\sigma_1^2}{\sigma_2^2} < 4.433 \times \frac{7.41^2}{5.67^2}) = (0.385 < \frac{\sigma_1^2}{\sigma_2^2} < 7.571);$$

$$\therefore 95\% \text{ confidence interval for } \frac{\sigma_1}{\sigma_2}: (0.62 < \frac{\sigma_1}{\sigma_2} < 2.75).$$

Since $1 \in \text{CI}$, there is no significant evidence against $\sigma_1 = \sigma_2$ in these data.

ii. $s^2 = \frac{1}{2}(s_1^2 + s_2^2)$, since the sample sizes are equal;

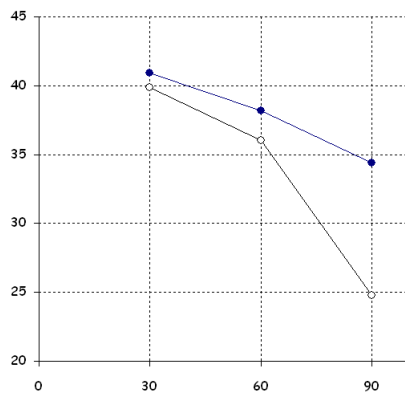
$$\text{hence } s = \sqrt{\frac{1}{2}(7.41^2 + 5.47^2)} = \sqrt{43.53} = 6.60.$$

- iii. $\frac{16S^2}{\sigma^2} \stackrel{d}{=} \chi_{16}^2 \Rightarrow$
 95% confidence interval for σ : $\left(\sqrt{\frac{16 \times 43.53}{28.85}}, \sqrt{\frac{16 \times 43.53}{5.908}} \right) = (4.91, 10.04)$
- iv. 95% confidence interval for $\mu_1 - \mu_2$: $7.4 \pm 2.120 \sqrt{43.53 \left(\frac{1}{9} + \frac{1}{9} \right)}$
 $= (7.4 \pm 6.59) = (0.81, 13.99).$

8. (a) i. error MS = $av(S^2) = \frac{1}{3}(20 + 24 + 31) = 25$

treatment	2	200	100	4
error	15	375	(25)	
total	17	(575)		

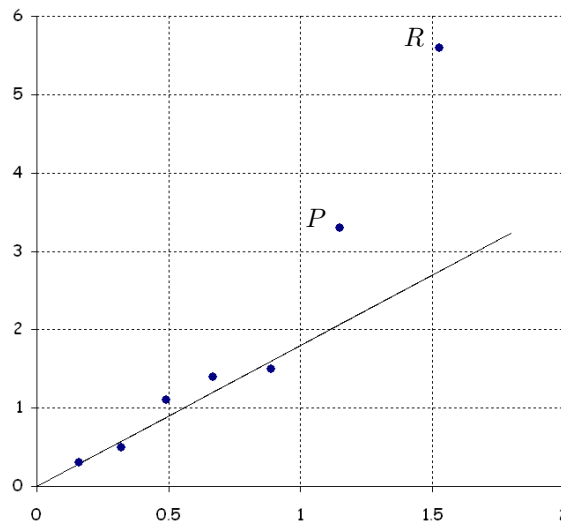
- ii. $H_0 \Rightarrow F \stackrel{d}{=} F_{2,15}, c = c_{0.95}(F_{2,15}) = 3.7;$
 $F = 4,$ and since $F > c,$ we reject $H_0.$ There is evidence of significant treatment effects. (It appears that both A and B are significantly greater than the control.)
- iii. $c_{0.95}(F_{1,6}) = 5.99, c_{0.95}(F_{2,6}) = 5.14;$ so all effects are significant: type, humidity and interaction are all significant. The cell mean plot is shown below:



Humidity reduces the required force; A is better than B (greater required force); humidity has a greater effect on B (it reduces the required force by more).

9. (a) The “full” model is fitted, which has as many parameters as degrees of freedom and therefore gives a “perfect” fit, with no degrees of freedom left over for error, and zero error, so that error SS = 0.
- (b) RMS = 3.3, 1.5, 5.6, 1.4, 0.5, 1.1, 0.3;

ordered RMS = 0.3, 0.5, 1.1, 1.4, 1.5, 3.3, 5.6;
 half-normal quantile = 0.16, 0.32, 0.49, 0.67, 0.89, 1.15, 1.53.



It appears that P and R are significant, and the others are not.

- (c) Assume that the effects other than P and R are zero, so that their sums of squares can be attributed to error.

P	1	11.1	$F = 9.4^*$
R	1	31.2	$F = 26.4^*$
e	5	5.9	$s^2 = 1.18$
total	7	48.2	

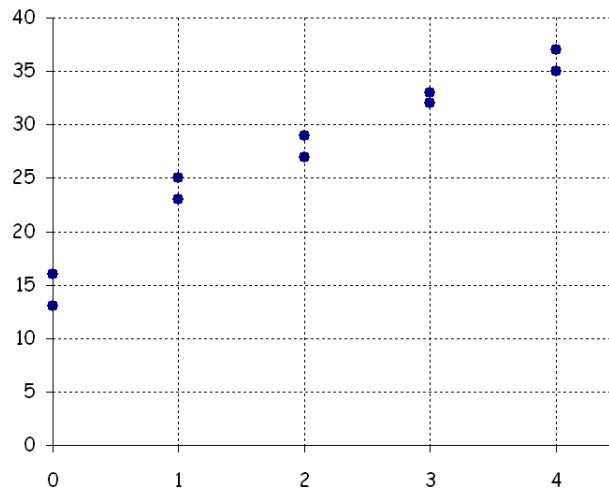
Since $c_{0.95}(F_{1,5}) = 6.61$, these results indicate that P and R are both significant.

- (d) $\hat{R} = -4.0$ ($= \text{av}(R1) - \text{av}(R0) = 13.9 - 17.9$);

$$95\% \text{ CI for } R: -4.0 \pm 2.571 \sqrt{1.18 \left(\frac{1}{4} + \frac{1}{4} \right)} = (-4.0 \pm 1.97) = (-5.97, -2.03).$$

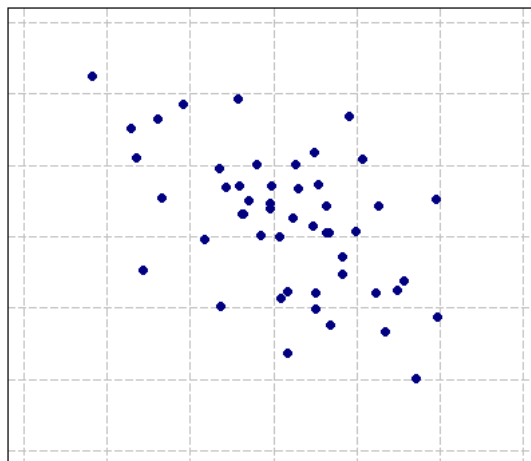
- (e) P has a significant (positive) effect; R has a significant (negative) effect; there is no evidence that Q has an effect, and nor is there evidence of any interactions.

10. (a) i. $\hat{\eta}_2 = \bar{y} = 27.0$, since $\hat{\eta}_x = \bar{y} + \hat{\beta}(x - \bar{x})$;
 95% confidence interval: $27.0 \pm 2.306 \sqrt{\frac{4.44}{10}} = (27.0 \pm 1.54) = (25.46, 28.54)$.
 ii. $\hat{\mu}_2 = \bar{y}_2 = 28.0$;
 95% confidence interval: $28.0 \pm 2.571 \sqrt{\frac{2.20}{2}} = (28.0 \pm 2.70) = (25.30, 30.70)$.



The straight-line regression assumption looks dubious: the observations at $x = 0$ don't seem to fit.

- (b) i. scatter plot with $\rho \approx -0.5$:



- ii. Table 20: $(-0.68 < \rho < -0.26)$.