

1. a) $\bar{y}_h = \frac{1}{n_h} \sum_{i=1}^{N_h} Y_{hi} - I_{hi}$ ✓

$$E \bar{y}_h = \frac{1}{n_h} \sum_{i=1}^{N_h} Y_{hi} - E I_{hi} ; \quad E I_{hi} = \frac{n_h}{N_h} \checkmark$$

$$= \frac{1}{N_h} \sum_{i=1}^{N_h} Y_{hi} = \mu_h \checkmark$$

b) $E \hat{\mu}_1 = \sum_{h=1}^L \frac{N_h}{N} \mu_h = \sum_{h=1}^L \sum_{i=1}^{N_h} Y_{hi} / N = \mu$ ✓

$E \hat{\mu}_2 \neq \mu$ in general (unless $\frac{n_h}{n} = \frac{N_h}{N}$) ✓

so prefer $\hat{\mu}_1$ as it is unbiased. ✓

c) use Lagrange multiplier

min $\text{Var } \hat{\mu}_1 = \sum_{h=1}^L W_h^2 S_h^2 (1-f_h) / n_h$ (from formulas) ✓

s.t. $\sum_{h=1}^L c_h n_h^2 = C$ ✓

equiv. to

min $\sum_h W_h^2 S_h^2 \left(\frac{1}{n_h} - \frac{1}{N_h} \right) + \lambda \left(\sum_h c_h n_h^2 - C \right)$ ✓

$$\frac{\partial}{\partial n_h} = - \frac{W_h^2 S_h^2}{n_h^2} + 2 \lambda c_h n_h$$

$$= 0 \implies \lambda = \frac{W_h^2 S_h^2}{2 c_h n_h^3} \checkmark$$

this is true for all h , so we have ✓

$$n_h^3 \propto \frac{W_h^2 S_h^2}{c_h}$$

2. Clearly a linear relationship between x & y , regression line probably doesn't go through the origin, so regression estimator appropriate ✓

$$\hat{\beta} = \Delta_{xy} / \Delta_x^2 = \frac{92.0}{47.7} = 1.93 \quad (3 \text{ sig figs}) \quad \checkmark$$

$$\begin{aligned} \hat{\mu}_{er} &= \bar{y} + \hat{\beta}(\mu_x - \bar{x}) \\ &= 215.1 + 1.93(100 - 97.8) \\ &= 219.3 \quad (3 \text{ sig figs}) \end{aligned}$$

$$\hat{\tau}_{er} = 21,930 \quad (\text{--- " ---}) \quad \checkmark$$

$$\begin{aligned} \widehat{\text{var}} \hat{\mu}_{er} &= \frac{1}{n} (1-f) \frac{h-1}{n-2} (\Delta_y^2 - \Delta_{xy}^2 / \Delta_x^2) \\ &= \frac{1}{10} \cdot \frac{9}{10} \cdot \frac{9}{8} (187.2 - (92.0)^2 / 47.7) \\ &= \frac{81}{800} \times 9.76 = 0.988 \quad (3 \text{ sig figs}) \quad \checkmark \end{aligned}$$

95% CI for τ

$$\begin{aligned} \hat{\tau}_{er} \pm t_{g(.975)} 100 \sqrt{0.988} \\ &= 21,930 \pm 2.306 \times 99.3 \\ &= 21,930 \pm 229 \quad (3 \text{ sig figs}) \quad \checkmark \end{aligned}$$

aliter using ratio estimator (max 5/6)

$$\hat{\tau}_{ratio} = 100 \hat{\mu}_{ratio} = 100 \frac{\bar{y}}{\bar{x}} \mu_x = 100 \frac{215.1}{97.8} 100 = 21,994 \quad (5 \text{ sig figs}) \quad \checkmark$$

$$\begin{aligned} \widehat{\text{var}}(\hat{\mu}_{ratio}) &= \frac{1}{n} (1-f) (\Delta_y^2 - 2r \Delta_{xy} + r^2 \Delta_x^2) \\ &= \frac{1}{10} \frac{9}{10} (187.2 - 2 \times \frac{215.1}{97.8} \times 92.0 + (\frac{215.1}{97.8})^2 \times 47.7) \\ &= 1.193 \quad (4 \text{ sig figs}) \quad \checkmark \end{aligned}$$

95% CI for τ

$$\begin{aligned} \hat{\tau}_{ratio} \pm t_{g(.975)} 100 \sqrt{1.193} \\ &= 21,994 \pm 2.262 \times 109.2 \\ &= 21,994 \pm 247 \quad (3 \text{ sig figs}) \quad \checkmark \end{aligned}$$

more 2/6 for SRS

3. cluster sample: 10 clusters $n = 10, N = 20$

$$\text{Var } \hat{\mu}_{cl} = S_b^2 (1-f)/n = S_b^2 \cdot \frac{1}{20} \quad \checkmark$$

simple random sample: sample size 20
 $n = 20, N = 100$

$$\text{Var } \hat{\mu}_{srs} = S^2 (1-f)/n = S^2 \cdot \frac{1}{25} \quad \checkmark$$

total SS = within SS + between SS

$$99 S^2 = \sum_{i=1}^{20} 4 S_i^2 + \sum_{i=1}^{20} 5 (\mu_i - \mu)^2$$

$$= 80 \bar{S}^2 + 95 S_b^2 \quad \checkmark \checkmark$$

$$\text{Var } \hat{\mu}_{cl} > \text{Var } \hat{\mu}_{srs}$$

$$\Leftrightarrow S_b^2 \cdot \frac{1}{20} > S^2 \cdot \frac{1}{25} \quad \checkmark$$

$$\Leftrightarrow \frac{99 S^2 - 80 \bar{S}^2}{95} \cdot \frac{1}{20} > S^2 \cdot \frac{1}{25} \quad \checkmark$$

$$\Leftrightarrow 99 \cdot 25 S^2 - 80 \cdot 25 \bar{S}^2 > 95 \cdot 20 S^2$$

$$\Leftrightarrow 575 S^2 > 2000 \bar{S}^2$$

$$\Leftrightarrow S^2 > 3.48 \bar{S}^2 \quad (3 \text{ sig figs}) \quad \checkmark \checkmark$$

prefer cluster sample when $S^2 < 3.48 \bar{S}^2$

4. a)
$$\begin{aligned} \mathbb{E} \hat{\mu} &= \mathbb{E} \mathbb{E}(\hat{\mu} \mid \hat{\mu}_1, \dots, \hat{\mu}_{10}) \quad \checkmark \\ &= \mathbb{E} \mathbb{E}\left(\frac{\sum_{c=1}^5 \hat{\mu}_{k(c)}/5}{\mid \hat{\mu}_1, \dots, \hat{\mu}_{10}}\right) \\ &= \mathbb{E} \mathbb{E}\left(\frac{\sum_{j=1}^{10} I_j \hat{\mu}_j}{5} \mid \hat{\mu}_1, \dots, \hat{\mu}_{10}\right) \quad \checkmark \\ &\quad \text{where } I_j = \begin{cases} 1 & \text{if } j = k(i) \\ & \text{some } i \\ 0 & \text{o/w} \end{cases} \\ &= \mathbb{E} \frac{1}{10} \sum_{j=1}^{10} \hat{\mu}_j \quad \text{since } \mathbb{E} I_j = \frac{5}{10} \quad \checkmark \\ &= \frac{1}{10} \sum_{j=1}^{10} \mu_j \quad \text{since } \mathbb{E} \hat{\mu}_j = \mu_j \quad \checkmark \\ &= \mu \quad (\text{equal sized clusters}) \end{aligned}$$

b) original sample:
 clusters: $k(1), \dots, k(5)$
 within cluster $k(i)$: $y_{k(i), j}$, $j = 1, \dots, 10$

bootstrap sample
 clusters: $k^*(1), \dots, k^*(5)$ resampled from $\{k(1), \dots, k(5)\}$ \checkmark

for each $k^*(i)$
 resample $y_{k^*(i), j}^*$ from $\{y_{k^*(i), j}\}_{j=1}^{10}$ \checkmark
 bootstrap replicate $\hat{\mu}^* = \sum_{c=1}^5 \left(\sum_{j=1}^{10} y_{k^*(c), j}^* / 10 \right) / 5$

let $y_{(1)}^*, \dots, y_{(B)}^*$ be bootstrap samples \checkmark
 and $\hat{\mu}^{*(1)}, \dots, \hat{\mu}^{*(B)}$ bootstrap replicates

then
$$\widehat{\text{var}}_B \hat{\mu} = \frac{1}{B-1} \sum_{c=1}^B (\hat{\mu}^{*(c)} - \hat{\mu}^{*(\cdot)})^2 \quad \checkmark$$

5. i) $(1, 0, 0) (0, 1, 1)_{12}$ ✓

ii) $(1 - B - B^4 + B^5) X_t = (1 + \beta_1 B + \beta_4 B^4 + \beta_1 \beta_4 B^5) Z_t$
 $(1 - B)(1 - B^4) X_t = (1 + \beta_1 B)(1 + \beta_4 B^4) Z_t$

$(0, 1, 1) (0, 1, 1)_4$ ✓✓

iii) $(1 - B^{12})^2 X_t = (1 - B^{12})^2 Z_t$

$(0, 0, 0) (0, 2, 2)_{12}$ ✓✓✓

6. a) MA(1) process $\phi(B) = 1$, $\theta(B) = 1 + \beta B$ ✓✓

$\gamma(0) = \sigma^2(1 + \beta^2)$

$\gamma(1) = \sigma^2 \beta$ ✓✓ where $\sigma^2 = \text{Var } Z_t$

$\frac{\gamma(1)}{\gamma(0)} = \frac{\beta}{1 + \beta^2} = \frac{2}{5}$

$2\beta^2 - 5\beta + 2 = 0$

$\beta = \frac{5 \pm \sqrt{25 - 16}}{4} = 2, 1/2$ ✓

$\beta = 1/2$ gives invertible process ✓✓

b) AR(2) process $\theta(B) = 1$, $\phi(B) = 1 - \alpha_1 B - \alpha_2 B^2$ ✓✓

auxiliary eqn $\lambda^2 - \alpha_1 \lambda - \alpha_2 = (\lambda - 5/6)(\lambda - 1/6)$
 $= \lambda^2 - \lambda + 5/36$ ✓✓

so $\alpha_1 = +1$, $\alpha_2 = -5/36$ ✓

Using first two Yule-Walker eqns get $\alpha_1 = \frac{52}{135}$ $\alpha_2 = \frac{2}{135}$

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$$z_t = x_t - \frac{1}{2}x_{t-1} + z_{t-1}$$

t	x_t	R_t
0	0	0
1	8	8 ✓
2	-12	-8 ✓
3	-6	-8 ✓
4	21	16 ✓
5	$5\frac{1}{2}$	0 ✓

8 a)

t	1	2	3	4	5	6	7	8	
M_t				1.3	1.4	1.9	2.6	3.1	(1 d.p.)

$$\hat{x}_8 = 3.1$$

✓

b)

t	1	2	3	4	5	6	7	8	
F_t	1.4	1.4	1.87	1.35	1.43	1.87	2.27	2.55	✓✓
			1.33	1.31	1.51	2.48	2.98	3.13	
X_t	1.4	1.3	1.3	1.6	2.9	3.2	3.2		

$$\hat{x}_8 = 2.6 \quad (1 \text{ d.p.})$$

c)

MA SSE for periods 4-7

$$.3^2 + 1.5^2 + 1.3^2 + .6^2 = 4.4 \quad (1 \text{ d.p.})$$

Exp Smoother SSE for periods 4-7

$$.25^2 + 1.47^2 + 1.33^2 + .93^2 = 4.9 \quad (1 \text{ d.p.})$$

no prefer 3 p.t. moving average smoother.

✓✓

9 a) $\hat{F}(x) = \sum_{i=1}^n \mathbb{1}_{\{x \leq x_i\}}$ ✓

b) Assume $\theta = E(F)$ some fn t
 then $\hat{\theta} = E(\hat{F})$ ✓

c) $\theta = \max_F X = \inf \{x : F(x) = 1\}$ ✓

$\hat{\theta} = \inf \{x : \hat{F}(x) = 1\} = \max_i \{x_i\}$ ✓✓

d) $\hat{F} \rightarrow F$ uniformly with prob. 1 as $n \rightarrow \infty$
 so, if t smooth enough, $t(\hat{F}) \rightarrow t(F)$ ✓✓

also \hat{F} is non-parametric ML est of F ✓

$$se_F \hat{\theta} = \sqrt{E_F (S(X) - E_F S(X))^2}$$

10 a) ~~Let $X^*(1), \dots, X^*(B)$ be iid samples from F
 and $\hat{\theta}^*(i) = S(X^*(i))$~~

~~then $se_B \hat{\theta} = \sqrt{\frac{1}{B-1} \sum (\hat{\theta}^*(i) - \hat{\theta}^*(i))^2}$~~

b) i) ^{old version of} as per (a) with $B = 250$ ✓✓

ii) $se_B \hat{\theta} \approx se_F \hat{\theta} = E_F (S(X) - E_F S(X))^2$ ✓✓
 $\approx se_F \hat{\theta} = E_F (S(X) - E_F S(X))^2$

reduce first error by increasing B ✓
 reduce second error by increasing n . ✓

iii) max sample ≤ 10 iff all elts of
 sample ≤ 10
 prob. a single elt. ≤ 10 is 0.6
 whence

$(0.6)^{10} = 0.006$ (3 d.p.). ✓✓

$$11. a) \widehat{\text{bias}}_B \hat{\theta} = \hat{\theta}^{*(.)} - E(\hat{F}) \quad \checkmark$$

$$\bar{\theta} = 2 E(\hat{F}) - \hat{\theta}^{*(.)} \quad \checkmark$$

$$b) \text{Var}_F \bar{\theta} = 4 \text{Var}_F E(\hat{F}) + \text{Var}_F \hat{\theta}^{*(.)}$$

$$\approx 4 (\widehat{\text{se}}_B \hat{\theta})^2 + \text{Var}_F \hat{\theta}^{*(.)}$$

$$\approx (4 + \frac{1}{B}) (\widehat{\text{se}}_B \hat{\theta})^2 \quad \checkmark \checkmark$$

$\bar{\theta}$ should be unbiased (approx) \checkmark
 so $\text{MSE} \bar{\theta} = \text{Var} \bar{\theta}$

$$c) \text{MSE}_F \hat{\theta} = \text{Var}_F \hat{\theta} + (\widehat{\text{bias}}_F \hat{\theta})^2$$

$$\approx (\widehat{\text{se}}_B \hat{\theta})^2 + (\widehat{\text{bias}}_B \hat{\theta})^2 \quad \checkmark$$

$$\text{so } \text{MSE}_F \bar{\theta} - \text{MSE}_F \hat{\theta}$$

$$\approx (3 + \frac{1}{B}) (\widehat{\text{se}}_B \hat{\theta})^2 - (\widehat{\text{bias}}_B \hat{\theta})^2$$

$$< 0 \quad \text{when } \widehat{\text{bias}}_B \hat{\theta} > \sqrt{3 + 1/B} \widehat{\text{se}}_B \hat{\theta}$$

$$\approx \sqrt{3} \widehat{\text{se}}_B \hat{\theta}$$

for B large. $\checkmark \checkmark$